

Fuzzy Difference-in-Differences (de Chaisemartin & D'Haultfœuille, REStud 2018)

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Introduction

- Goal: study **instrumented difference-in-differences** with **heterogeneity in treatment effects**
- **Fuzzy DiD** is a commonly used design
 - But it is not always explicitly recognized as such in empirical work (e.g. Duflo 2001)
- In a related article, Blundell and Costa Dias (2009) discusses fuzzy DiD with heterogeneity

Today's Presentation

- ① Identification of a (conditional) LATE with a **standard** Wald estimand and **strong** assumptions
- ② Identification of a (conditional) LATE with a **corrected** Wald estimand and **other** assumptions
- ③ Extension to designs with **multiple treatment groups**

① Model Setup

② Identification Results

③ Extension to Multiple Treatment Groups

Model Setup

- $T \in \{0, 1\}$ denotes **time**, $G \in \{0, 1\}$ indicates a **time-invariant group**
- D is a **binary treatment**, **not** a **deterministic** function of G and T , i.e., $D \neq GT$
- $Y \in \mathbb{R}$ is an **outcome** of interest
- $Z \equiv GT$ is a **binary instrument** (not nested in Hudson, Hull, and Liebersohn 2017)
- Unlike sharp designs, **fuzzy** designs allow for
 - Units to be treated in the control group ($G = 0$)
 - Units to be treated (in either group) in $T = 0$

The IV-DiD Wald Estimand

- **Reduced-form** and **first-stage** (saturated) linear regressions:

$$Y = \alpha + \beta G + \gamma T + \delta GT + U$$

$$D = \lambda + \eta G + \phi T + \rho GT + V$$

- The coefficient associated with D in the structural equation is the **IV-DiD Wald estimand**:

$$\omega \equiv \frac{\Delta_Y(1) - \Delta_Y(0)}{\Delta_D(1) - \Delta_D(0)}$$

where $\Delta_Y(g)$ and $\Delta_D(g)$ denote **time trends** in group $g \in \{0, 1\}$

Further Assumptions

Without loss of generality, define D and G such that

- 1 The treatment rate **increases** over time in the **treated** group,

$$\mathbb{E}[D|G = 1, T = 1] - \mathbb{E}[D|G = 1, T = 0] > 0$$

- 2 The treatment rate in the **control** group **does not increase more** than in the **treated** group,

$$\mathbb{E}[D|G = 1, T = 1] - \mathbb{E}[D|G = 1, T = 0] > \mathbb{E}[D|G = 0, T = 1] - \mathbb{E}[D|G = 0, T = 0]$$

Potential Outcomes and Potential Treatments

- D and Y are linked by **potential outcomes** $Y(0), Y(1)$
- **Potential treatments** are $D(0), D(1)$, where $D = D(t)$ is observed (G is subsumed)
 - Both potential treatments are **observed** in a repeated cross section
 - Potential treatments are **independent of time within each group**
- Within each group, units **switch** treatment only in **one direction**:

$$\mathbb{P}(D(1) \geq D(0)|G) = 1 \quad \text{or} \quad \mathbb{P}(D(1) \leq D(0)|G) = 1$$

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Identification Result #1

- Further assume that
 - **Common trends** holds, i.e., $\mathbb{E}[Y(0)|G, T = 1] - \mathbb{E}[Y(0)|G, T = 0]$ does not depend on G
 - The **ATE** among units **treated in the pre-period** is **stable** over time **within** each group, i.e.,

$$\mathbb{E}[Y(1) - Y(0)|G, T = 1, D(0) = 1] = \mathbb{E}[Y(1) - Y(0)|G, T = 0, D(0) = 1]$$
- Then the IV-DiD Wald estimand identifies a **weighted average of two causal parameters**:
 - 1 $\tau_1 \equiv \mathbb{E}[Y(1) - Y(0)|G = 1, T = 1, D(1) > D(0)]$, the ATE among **treated “switchers”**
 - 2 $\tau_0 \equiv \mathbb{E}[Y(1) - Y(0)|G = 0, T = 1, D(1) \neq D(0)]$, the ATE among **control “switchers”**

Identification Result #1

- Case (a): the treatment rate **increases** in the control group. Then

$$\omega = \alpha\tau_1 + (1 - \alpha)\tau_0 \quad \text{with} \quad \alpha \equiv \frac{\mathbb{P}(D(1) > D(0)|G = 1)}{\mathbb{P}(D(1) > D(0)|G = 1) - \mathbb{P}(D(1) > D(0)|G = 0)}$$

$\alpha > 1 \implies$ the IV-DiD estimand **negatively weights** the ATE among **control switchers**

- Case (b): the treatment rate **decreases** in the control group. Then

$$\omega = \alpha\tau_1 + (1 - \alpha)\tau_0 \quad \text{with} \quad \alpha \equiv \frac{\mathbb{P}(D(1) > D(0)|G = 1)}{\mathbb{P}(D(1) > D(0)|G = 1) + \mathbb{P}(D(1) < D(0)|G = 0)}$$

$\alpha \in (0, 1) \implies$ the IV-DiD estimand is a **convex combination** of **ATEs among switchers**

Identification Result #2

- **Modify** the assumptions made for identification result #1
- Instead, assume some version of **conditional common trends**:

$$\mathbb{E}[Y(d)|G, T = 1, D(0) = d] - \mathbb{E}[Y(d)|G, T = 0, D(0) = d]$$

does not depend on G for $d \in \{0, 1\}$

- Consider a **“time-corrected” Wald estimand**:

$$\omega_{TC} \equiv \frac{\mathbb{E}[Y|G = 1, T = 1] - \mathbb{E}[Y + (1 - D)\delta_0 + D\delta_1|G = 1, T = 0]}{\mathbb{E}[D|G = 1, T = 1] - \mathbb{E}[D|G = 1, T = 0]}$$

with

$$\delta_d \equiv \mathbb{E}[Y|D = d, G = 0, T = 1] - \mathbb{E}[Y|D = d, G = 0, T = 0] \quad \text{for } d \in \{0, 1\}$$

Identification Result #2

- ω_{TC} identifies τ_1 , the ATE among **switchers in the treated group**
- Intuition behind this “correction”:
 - Begin with $\mathbb{E}[Y|G = 1, T = 0]$
 - Add $\delta_0 \times \mathbb{P}(D = 0|G = 1, T = 0)$ and $\delta_1 \times \mathbb{P}(D = 1|G = 1, T = 0)$
 - Obtain a counterfactual $\mathbb{E}[Y|G = 1, T = 1]$ purged from the contribution of switchers
 - Contrast between $\mathbb{E}[Y|G = 1, T = 1]$ and “corrected” $\mathbb{E}[Y|G = 1, T = 0]$ \rightarrow switchers
 - Scale the numerator by the evolution of the treatment rate in the treated group

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Extension to Multiple Treatment Groups

These identification results can be extended to situations with multiple treatment groups

- **Groups** $G \in \{0, 1, \dots, \bar{g}\}$
- Partition them into **three super-groups** based on how the treatment rate evolves
 - Group g belongs to \mathcal{G}_i , \mathcal{G}_s , or \mathcal{G}_d if $\Delta_D(g)$ increases, is stable, or decreases
- The target parameter becomes the **ATE among all switchers**, i.e.,

$$\tau^* \equiv \mathbb{E} \left[Y(1) - Y(0) \middle| T = 1, \bigcup_{g=0}^{\bar{g}} \{D(0) \neq D(1), G = g\} \right]$$

- For compactness of notation, define $G^* \equiv \mathbb{I}[G \in \mathcal{G}_i] - \mathbb{I}[G \in \mathcal{G}_d] \in \{-1, 0, 1\}$

Extension to Multiple Treatment Groups

Under the same assumptions as in **identification result #1**, τ^* is identified as follows:

- 1 Compute **four difference-in-differences contrasts**:

$$\text{DiD}_R^*(g, g') \equiv \Delta_R(g) - \Delta_R(g')$$

where R is either Y or D and $(g, g') \in \{(1, 0), (0, -1)\}$

- 2 Compute two Wald estimands by taking **ratios of DiD contrasts**:

$$\omega_{\text{DiD}}^*(1, 0) \equiv \frac{\text{DiD}_Y^*(1, 0)}{\text{DiD}_D^*(1, 0)} \quad \omega_{\text{DiD}}^*(0, -1) \equiv \frac{\text{DiD}_Y^*(0, -1)}{\text{DiD}_D^*(0, -1)}$$

- 3 Compute a **convex combination** of these two Wald estimands:

$$\omega_{\text{DiD}}^* \equiv \theta \omega_{\text{DiD}}^*(1, 0) + (1 - \theta) \omega_{\text{DiD}}^*(0, -1)$$

Extension to Multiple Treatment Groups

Under the same assumption as in **identification result #2**, τ^* is identified as follows:

- 1 Compute **two time correction terms** (for $d \in \{0, 1\}$):

$$\delta_d^* \equiv \mathbb{E}[Y|D = d, G^* = 0, T = 1] - \mathbb{E}[Y|D = d, G^* = 0, T = 0]$$

- 2 Compute **two time-corrected Wald ratios** (for $g \in \{-1, 1\}$):

$$\omega_{TC}^*(g) \equiv \frac{\mathbb{E}[Y|G^* = g, T = 1] - \mathbb{E}[Y + (1 - D)\delta_0^* + D\delta_1^*|G^* = g, T = 0]}{\mathbb{E}[D|G^* = g, T = 1] - \mathbb{E}[D|G^* = g, T = 0]}$$

- 3 Compute a **convex combination** of these two time-corrected Wald ratios:

$$\omega_{TC}^* \equiv \theta \omega_{TC}^*(1) + (1 - \theta) \omega_{TC}^*(-1)$$

Extension to Multiple Treatment Groups

- In practice, super-groups $G^* \in \{-1, 0, 1\}$ may be **known** ex ante or need to be **estimated**
 - Estimation is likely necessary when the treatment varies at the unit level
- The authors develop a **data-based procedure** to classify groups into three super-groups
 - This hinges on running t-tests within each group to compare the treatment rate over time
- Once super-groups have been determined, target parameters can be estimated