

“Difference-in-Differences with multiple time periods” Callaway & Sant’Anna (JoE, 2021)

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April 26, 2022

Big Picture

- extensions to canonical diff-in-diff model
 - multiple time periods
 - variation in treatment timing
 - treatment effect heterogeneity
 - “parallel trends assumption” only conditional on covariates
- standpoint
 - TWFE can be weird/hard to interpret in these cases
 - What else could/should we do instead?
- plan
 - 1 identify disaggregated “group-time ATEs”
 - 2 aggregation of these to causal effects of interest
 - 3 estimation
 - outcome regression
 - inverse probability weighting
 - doubly-robust
 - 4 inference
 - bootstrap
 - joint inference

Model and Notation

- time periods $t \in \{1, \dots, \bar{t}\}$
- binary treatment $D_{i,t} \in \{0, 1\}$
- all units are untreated at the start, i.e. $D_{i,1} = 0$
- treatment is irreversible, i.e. $D_{i,t+1} \geq D_{i,t}$ for all $t = 1, \dots, \bar{t} - 1$
- $C_i \in \{0, 1\}$ is 1 iff i is never treated
- G_i is the time period in which i is first treated ($G_i = \infty$ if $C_i = 1$)
- hence the entire treatment path is characterized by G_i
- Let $G_{i,g} \in \{0, 1\}$ be 1 iff i is first treated in period g
- potential outcomes

$$Y_{i,t} = Y_{i,t}(0) + \sum_{g=2}^{\bar{t}} (Y_{i,t}(g) - Y_{i,t}(0)) G_{i,g}$$

Group-Time Average Treatment Effects

First objective is to identify

$$ATT(g, t) := \mathbb{E}[Y_{i,t}(g) - Y_{i,t}(0) | G_{i,g} = 1]$$

for all $g, t \in \{2, \dots, \bar{t}\}$.

Assumptions I

- 1 treatment
 - all units start untreated
 - treatment is irreversible
- 2 panel random sampling
 - $(Y_{i,1}, \dots, Y_{i,\bar{t}}, X_i, D_{i,1}, \dots, D_{i,\bar{t}})$ is i.i.d. over i
- 3 limited treatment anticipation: there is a known $\delta > 0$ such that for all $g < \max\{G\}$, $t \in \{1, \dots, \bar{t}\}$ for which $t < g - \delta$

$$\mathbb{E}[Y_{i,t}(g)|X_i, G_{i,g} = 1] = \mathbb{E}[Y_{i,t}(0)|X_i, G_{i,g} = 1]$$

Assumptions II

- 4 conditional parallel trends based on a 'never-treated' group
 Let δ be the anticipation horizon from assumption 3. For each $g < \max\{G\}$, $t \in \{2, \dots, \bar{t}\}$ such that $t \geq g - \delta$,
- $$\mathbb{E}[Y_{i,t}(0) - Y_{i,t-1}(0)|X_i, G_{i,g} = 1] = \mathbb{E}[Y_{i,t}(0) - Y_{i,t-1}(0)|X_i, C_i = 1].$$
- 5 conditional parallel trends based on a 'not-yet' group
 Let δ be the anticipation horizon from assumption 3. For each $g < \max\{G\}$, $t \in \{2, \dots, \bar{t}\}$ and each $s, t \in (2, \dots, \bar{t})$ such that $t \geq g - \delta$ and $t + \delta \leq s < \max\{G\}$
- $$\begin{aligned} &\mathbb{E}[Y_{i,t}(0) - Y_{i,t-1}(0)|X_i, G_{i,g} = 1] = \\ &\mathbb{E}[Y_{i,t}(0) - Y_{i,t-1}(0)|X_i, D_{i,s} = 0, G_{i,g} = 0]. \end{aligned}$$
- 6 For each $t \geq 2$, $g < \max\{G\}$, there exists some $\varepsilon > 0$ such that $\mathbb{P}[G_g = 1] > \varepsilon$ and
- $$p_{g,t}(X) := \mathbb{P}[G_g = 1|X, G_g + (1 - D_t)(1 - G_g) = 1] < 1 - \varepsilon.$$

Identification Result

Theorem 1

Suppose Assumption 1-4 and 6 hold. Then for all g and t such that $2 + \delta \leq g < \max\{G\}$ and $t \geq g - \delta$

$$ATT(g, t) = \mathbb{E} \left[\left(\frac{G_g}{\mathbb{E}[G_g]} - \frac{\frac{\rho_g(X)C}{1-\rho_g(X)}}{\mathbb{E} \left[\frac{\rho_g(X)C}{1-\rho_g(X)} \right]} \right) (Y_t - Y_{g-\delta-1} - m_{g,t,\delta}(X)) \right],$$

where

$$m_{g,t,\delta}(X) = \mathbb{E}[Y_t - Y_{g-\delta-1} | X, C = 1].$$

Choosing the “Right” Comparison Group

- A4 (\exists never-treated gr.) can be replaced w/ A5 (\exists “not-yet” gr.)
- equation in the thm changes slightly, see paper
- both condition on X , see eg [HIT97] for motivation
- If never-treated *and* not-yet treated gr exist, which one to use?
 - Callaway & Sant’Anna favor A4 in applications if
 - never-treated group is large enough
 - never-treated group is “similar enough” to treated groups
 - using a not-yet treated units with A5
 - allows larger comparison groups
 - restricts pre-(anticipation)treatment trends of “not-yet” treated groups
 - can be problematic, see eg [MS21]
 - pretesting to select A4/A5 is problematic, see [Rot19]

Target Parameter

Callaway and Sant'Anna consider linear aggregators of the form

$$\theta = \sum_g \sum_{t=2}^{\bar{t}} w(g, t) ATT(g, t)$$

for some “carefully-chosen (known or estimable) weighting functions specified by the researcher such that θ can be used to address a well-posed empirical/policy question.”

In particular,

- $w(g, t) \geq 0$ for all g and $t \geq 2$,
- $\sum_g \sum_{t=2}^{\bar{t}} w(g, t) = 1$.

Examples

- What are “overall” treatment effects?
- How does the effect of participating in the treatment vary with length of exposure to the treatment?
- Do groups that are treated earlier have, on average, higher/lower average treatment effects than groups that are treated later?
- What is the cumulative average treatment effect of the policy across all groups until some particular point in time?
 - Assume $\delta = 0$.
 - Then Callaway & Sant’Anna propose to consider

$$\theta^{\text{cumulative}} = \sum_{g < \max G} \mathbf{1}_{t \geq g} \mathbb{P}[G = g | G < t] ATT(g, t).$$

Estimation

- recall the identification result

$$ATT(g, t) = \mathbb{E} \left[\left(\frac{G_g}{\mathbb{E}[G_g]} - \frac{\frac{p_g(X)C}{1-p_g(X)}}{\mathbb{E} \left[\frac{p_g(X)C}{1-p_g(X)} \right]} \right) (Y_t - Y_{g-\delta-1} - m_{g,t,\delta}(X)) \right]$$

- procedure

- estimate $ATT(g, t)$
 - estimate $p_g(X)$ and $m_{g,t,\delta}(X)$
 - use sample analog of \mathbb{E} above, plugging in $\hat{p}_g(X)$ and $\hat{m}_{g,t,\delta}(X)$
- calculate/estimate weights $w(g, t)$
- calculate θ

Estimation

- recall the identification result

$$ATT(g, t) = \mathbb{E} \left[\left(\frac{G_g}{\mathbb{E}[G_g]} - \frac{\frac{p_g(X)C}{1-p_g(X)}}{\mathbb{E} \left[\frac{p_g(X)C}{1-p_g(X)} \right]} \right) (Y_t - Y_{g-\delta-1} - m_{g,t,\delta}(X)) \right]$$

- the identification result in the paper is for three estimators
 - outcome regression
 - leaves out blue term
 - requires *correct* model for $m_{g,t,\delta}(X)$ (outcome evol of comp group)
 - inverse probability weighting
 - leaves out red term
 - requires *correct* model for $p_g(X)$ (propensity score)
 - doubly robust
 - includes both terms
 - robust to misspecification of either $m_{g,t,\delta}(X)$ or $p_g(X)$

Inference: Setting and Assumptions

- asymptotic regime
 - \bar{t} fixed
 - n goes to ∞
- assumptions
 - ⑦ parametric model for $m_{g,t,\delta}(X)$ or $p_g(X)$
 - sufficiently smooth
 - $\exists \sqrt{n}$ strongly consistent estimator for parameters
 - ⑧ param. model for at least one of $m_{g,t,\delta}(X)$ or $p_g(X)$ is correct
 - ⑨ integrability assumptions
- allows different estimators ((non-)least squares, mle,...)

Inference: Theorem

Theorem 2

Suppose Assumption 1-4 and 6-9 hold. Then for all g and t such that $2 + \delta \leq g < \max G$ and $g - \delta \leq t \leq \bar{t} - \delta$

$$\sqrt{n} \left(A\hat{T}T_{t \geq g - \delta} - ATT_{t \geq g - \delta} \right) = \frac{1}{\sqrt{n}} \sum_{i=1}^n \Psi_i + o_p(1)$$

where an analytic expression for Ψ_i is provided in the paper. Furthermore

$$\sqrt{n} \left(A\hat{T}T_{t \geq g - \delta} - ATT_{t \geq g - \delta} \right) \xrightarrow{d} \mathcal{N}(0, \Sigma)$$

where an analytic expression for Σ is provided in the paper.

Inference: Implementation

- could use analytic expression to estimate variance, Slutsky etc
- Callaway & Sant'Anna prefer the multiplier bootstrap
 - does not require re-estimation of propensity score in each bootstrap iteration
 - simplifies simultaneous inference
 - for each bootstrap iteration $s = 1, \dots, S$
 - 1 draw a sample of N iid random variables with zero mean, unit variance, and finite third moment, independent of the original data, eg

$$V = \begin{cases} 1 - \kappa & \text{w.p. } \frac{\kappa}{\sqrt{5}}, \\ \kappa & \text{w.p. } 1 - \frac{\kappa}{\sqrt{5}} \end{cases}$$

for $\kappa = (\sqrt{5} + 1)/2$.

- 2 $A\hat{T}T^s(t, g) = A\hat{T}T + \overline{V\hat{\Psi}}$.

- this yields valid point-wise inference
- see paper for formulas for simultaneous adjustments
- can draw cluster-specific V 's for cluster inference (if clusters are large)

Lemma for Proof of Theorem 1

$$\begin{aligned}
 & \mathbb{E}[Y_t(g) - Y_t(0)|X, G_g = 1] \\
 = & \mathbb{E}[Y_t(g) - Y_{g-\delta-1}(0)|X, G_g = 1] - \mathbb{E}[Y_t(0) - Y_{g-\delta-1}(0)|X, G_g = 1] \\
 = & \mathbb{E}[Y_t(g) - Y_{g-\delta-1}(0)|X, G_g = 1] - \sum_{l=0}^{t-g-\delta} \mathbb{E}[Y_{t-l}(0) - Y_{t-l-l}(0)|X, G_g = 1] \\
 \stackrel{4}{=} & \mathbb{E}[Y_t(g) - Y_{g-\delta-1}(0)|X, G_g = 1] - \sum_{l=0}^{t-g-\delta} \mathbb{E}[Y_{t-l}(0) - Y_{t-l-l}(0)|X, C = 1] \\
 = & \mathbb{E}[Y_t(g) - Y_{g-\delta-1}(0)|X, G_g = 1] - \mathbb{E}[Y_t(0) - Y_{g-\delta-1}(0)|X, C = 1] \\
 \stackrel{3}{=} & \mathbb{E}[Y_t - Y_{g-\delta-1}|X, G_g = 1] - \mathbb{E}[Y_t - Y_{g-\delta-1}|X, C = 1].
 \end{aligned}$$

Note that assumption 1 is used whenever we write the potential outcomes as function of g .

Proof of Theorem 1 for the Outcome Regression

$$\begin{aligned}
 & ATT(g, t) \\
 &= \mathbb{E}[\mathbb{E}[Y_t(g) - Y_t(0) | X, G_g = 1] | G_g = 1] \\
 &= \mathbb{E}[\mathbb{E}[Y_t - Y_{g-\delta-1} | X, G_g = 1] - \mathbb{E}[Y_t - Y_{g-\delta-1} | X, C = 1] | G_g = 1] \\
 &= \mathbb{E}[Y_t - Y_{g-\delta-1} | G_g = 1] - \underbrace{\mathbb{E}[\mathbb{E}[Y_t - Y_{g-\delta-1} | X, C = 1] | G_g = 1]}_{=: m_{g,t,\delta}(X)} \\
 &= \mathbb{E}[Y_t - Y_{g-\delta-1} - m_{g,t,\delta}(X) | G_g = 1] \\
 &= \mathbb{E} \left[\frac{G_g}{\mathbb{P}[G_g = 1]} (Y_t - Y_{g-\delta-1} - m_{g,t,\delta}(X)) \right] \\
 &= \mathbb{E} \left[\frac{G_g}{\mathbb{E}[G_g]} (Y_t - Y_{g-\delta-1} - m_{g,t,\delta}(X)) \right].
 \end{aligned}$$

References I



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Michelle Marcus and Pedro HC Sant'Anna, *The role of parallel trends in event study settings: An application to environmental economics*, *Journal of the Association of Environmental and Resource Economists* **8** (2021), no. 2, 235–275.



Jonathan Roth, *Pre-test with caution: Event-study estimates after testing for parallel trends*, Department of Economics, Harvard University, Unpublished manuscript (2019).